# Overfitting, Bias / Variance Analysis 

Professor Ameet Talwalkar

Slide Credit: Professor Fei Sha

## Outline

(1) Administration

(2) Review of last lecture
(3) Basic ideas to overcome overfitting
4) Bias/Variance Analysis

## Announcements

- HW2 (and remaining HW1) will be returned today after class (see Nikos and Amogh)
- HW3 and HW4 due on Thursday
- No new HW assigned on Thursday (to give you time to work on project proposal)
- Project proposal guideline posted on course website last week


## CS260: Project Proposal

## Professor Ameet Talwalkar

$$
\text { Due: } 11 / 5 / 15
$$

Project proposals should be $1-2$ pages long and contain the following information:

- Motivation: What high-level problem are you studying? Why is it important?
- Background: What previous work exists on your topic? How does your proposed project differ from these existing works? Note: If you have prior experience working on this problem, please clearly describe how your proposed work for this class project differs from your prior work.
- Proposed Work: What do you plan to do? What methods will you use and why will you use them?
- Timeline: What is your timeline for performing this work?
- Deliverables / Evaluation: What is the expected outcome of your work? How will you evaluate the quality of your work?
- Data: If your project involves data, what data are you using? How did you obtain this data? Why is this data interesting / relevant to your problem? What computing resources will you be using to analyze this data?
- Software Tools / Libraries: What (if any) software will be used for this project? Please describe any third-party software libraries you plan to use.
- Team: Who is working on this project (at most two students can work on a project)?
- Prior Discussion: You are required to discuss this project with Professor Talwalkar or one of the TAs prior to submitting your proposal. Please list who you spoke with and when. Note: You are also encouraged to speak with us as you continue working on the project if you have questions.


## Outline

## (1) Administration

(2) Review of last lecture

- Linear Regression
- Ridge Regression for Numerical Purposes
- Non-linear Basis
(3) Basic ideas to overcome overfitting

4) Bias/Variance Analysis

## Linear regression 1D

## Setup

- Input: $\boldsymbol{x} \in \mathbb{R}^{D}$ (covariates, predictors, features, etc)
- Output: $y \in \mathbb{R}$ (responses, targets, outcomes, outputs, etc)
- Training data: $\mathcal{D}=\left\{\left(\boldsymbol{x}_{n}, y_{n}\right), n=1,2, \ldots, \mathrm{~N}\right\}$
- Model: $f: \boldsymbol{x} \rightarrow y$, with $f(\boldsymbol{x})=w_{0}+\sum_{d} w_{d} x_{d}=w_{0}+\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}$

We also sometimes call $\tilde{\boldsymbol{w}}=\left[\begin{array}{lllll}w_{0} & w_{1} & w_{2} & \cdots & w_{\mathrm{D}}\end{array}\right]^{\mathrm{T}}$ parameters too!

## Linear regression 1D

## Setup

- Input: $\boldsymbol{x} \in \mathbb{R}^{D}$ (covariates, predictors, features, etc)
- Output: $y \in \mathbb{R}$ (responses, targets, outcomes, outputs, etc)
- Training data: $\mathcal{D}=\left\{\left(\boldsymbol{x}_{n}, y_{n}\right), n=1,2, \ldots, \mathrm{~N}\right\}$
- Model: $f: \boldsymbol{x} \rightarrow y$, with $f(\boldsymbol{x})=w_{0}+\sum_{d} w_{d} x_{d}=w_{0}+\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}$

We also sometimes call $\tilde{\boldsymbol{w}}=\left[\begin{array}{lllll}w_{0} & w_{1} & w_{2} & \cdots & w_{\mathrm{D}}\end{array}\right]^{\mathrm{T}}$ parameters too!
Least Mean Squares (LMS) Objective: Minimize squared difference on training data (or residual sum of squares)

$$
R S S(\tilde{\boldsymbol{w}})=\sum_{n}\left[y_{n}-f\left(\boldsymbol{x}_{n}\right)\right]^{2}=\sum_{n}\left[y_{n}-\left(w_{0}+\sum_{d} w_{d} x_{n d}\right)\right]^{2}
$$

1D Solution: Identify stationary points by taking derivative with respect to parameters and setting to zero, yielding 'normal equations'

## LMS when $\boldsymbol{x}$ is D-dimensional

$R S S(\tilde{\boldsymbol{w}})$ in matrix form

$$
R S S(\tilde{\boldsymbol{w}})=\sum_{n}\left[y_{n}-\left(w_{0}+\sum_{d} w_{d} x_{n d}\right)\right]^{2}=\sum_{n}\left[y_{n}-\tilde{\boldsymbol{w}}^{\mathrm{T}} \tilde{\boldsymbol{x}}_{n}\right]^{2}
$$

where we have redefined some variables (by augmenting)

$$
\tilde{\boldsymbol{x}} \leftarrow\left[\begin{array}{lllll}
1 & x_{1} & x_{2} & \ldots & x_{\mathrm{D}}
\end{array}\right]^{\mathrm{T}}, \quad \tilde{\boldsymbol{w}} \leftarrow\left[\begin{array}{lllll}
w_{0} & w_{1} & w_{2} & \ldots & w_{\mathrm{D}}
\end{array}\right]^{\mathrm{T}}
$$

## LMS when $\boldsymbol{x}$ is D-dimensional

$R S S(\tilde{\boldsymbol{w}})$ in matrix form

$$
R S S(\tilde{\boldsymbol{w}})=\sum_{n}\left[y_{n}-\left(w_{0}+\sum_{d} w_{d} x_{n d}\right)\right]^{2}=\sum_{n}\left[y_{n}-\tilde{\boldsymbol{w}}^{\mathrm{T}} \tilde{\boldsymbol{x}}_{n}\right]^{2}
$$

where we have redefined some variables (by augmenting)

$$
\tilde{\boldsymbol{x}} \leftarrow\left[\begin{array}{lllll}
1 & x_{1} & x_{2} & \ldots & x_{\mathrm{D}}
\end{array}\right]^{\mathrm{T}}, \quad \tilde{\boldsymbol{w}} \leftarrow\left[\begin{array}{lllll}
w_{0} & w_{1} & w_{2} & \ldots & w_{\mathrm{D}}
\end{array}\right]^{\mathrm{T}}
$$

which leads to

$$
\begin{aligned}
R S S(\tilde{\boldsymbol{w}}) & =\sum_{n}\left(y_{n}-\tilde{\boldsymbol{w}}^{\mathrm{T}} \tilde{\boldsymbol{x}}_{n}\right)\left(y_{n}-\tilde{\boldsymbol{x}}_{n}^{\mathrm{T}} \tilde{\boldsymbol{w}}\right) \\
& =\sum_{n} \tilde{\boldsymbol{w}}^{\mathrm{T}} \tilde{\boldsymbol{x}}_{n} \tilde{\boldsymbol{x}}_{n}^{\mathrm{T}} \tilde{\boldsymbol{w}}-2 y_{n} \tilde{\boldsymbol{x}}_{n}^{\mathrm{T}} \tilde{\boldsymbol{w}}+\text { const. } \\
& =\left\{\tilde{\boldsymbol{w}}^{\mathrm{T}}\left(\sum_{n} \tilde{\boldsymbol{x}}_{n} \tilde{\boldsymbol{x}}_{n}^{\mathrm{T}}\right) \tilde{\boldsymbol{w}}-2\left(\sum_{n} y_{n} \tilde{\boldsymbol{x}}_{n}^{\mathrm{T}}\right) \tilde{\boldsymbol{w}}\right\}+\text { const. }
\end{aligned}
$$

$R S S(\tilde{\boldsymbol{w}})$ in new notation

Design matrix and target vector

$$
\tilde{\boldsymbol{X}}=\left(\begin{array}{c}
\tilde{\boldsymbol{x}}_{1}^{\mathrm{T}} \\
\tilde{\boldsymbol{x}}_{2}^{\mathrm{T}} \\
\vdots \\
\tilde{\boldsymbol{x}}_{\mathrm{N}}^{\mathrm{T}}
\end{array}\right) \in \mathbb{R}^{\mathrm{N} \times(D+1)}, \quad \boldsymbol{y}=\left(\begin{array}{c}
y_{1} \\
y_{2} \\
\vdots \\
y_{\mathrm{N}}
\end{array}\right)
$$

Compact expression

$$
R S S(\tilde{\boldsymbol{w}})=\|\tilde{\boldsymbol{X}} \tilde{\boldsymbol{w}}-\boldsymbol{y}\|_{2}^{2}=\left\{\tilde{\boldsymbol{w}}^{\mathrm{T}} \tilde{\boldsymbol{X}}^{\mathrm{T}} \tilde{\boldsymbol{X}} \tilde{\boldsymbol{w}}-2\left(\tilde{\boldsymbol{X}}^{\mathrm{T}} \boldsymbol{y}\right)^{\mathrm{T}} \tilde{\boldsymbol{w}}\right\}+\text { const }
$$

## Solution in matrix form

Compact expression

$$
R S S(\tilde{\boldsymbol{w}})=\|\tilde{\boldsymbol{X}} \tilde{\boldsymbol{w}}-\boldsymbol{y}\|_{2}^{2}=\left\{\tilde{\boldsymbol{w}}^{\mathrm{T}} \tilde{\boldsymbol{X}}^{\mathrm{T}} \tilde{\boldsymbol{X}} \tilde{\boldsymbol{w}}-2\left(\tilde{\boldsymbol{X}}^{\mathrm{T}} \boldsymbol{y}\right)^{\mathrm{T}} \tilde{\boldsymbol{w}}\right\}+\text { const }
$$

Gradients of Linear and Quadratic Functions

- $\nabla_{x} \boldsymbol{b}^{\top} \boldsymbol{x}=\boldsymbol{b}$
- $\nabla_{\boldsymbol{x}} \boldsymbol{x}^{\top} \boldsymbol{A} \boldsymbol{x}=2 \boldsymbol{A} \boldsymbol{x}$ (symmetric $\boldsymbol{A}$ )


## Normal equation

$$
\nabla_{\tilde{\boldsymbol{w}}} R S S(\tilde{\boldsymbol{w}}) \propto \tilde{\boldsymbol{X}}^{\mathrm{T}} \tilde{\boldsymbol{X}} \boldsymbol{w}-\tilde{\boldsymbol{X}}^{\mathrm{T}} \boldsymbol{y}=0
$$

This leads to the least-mean-square (LMS) solution

$$
\tilde{\boldsymbol{w}}^{L M S}=\left(\tilde{\boldsymbol{X}}^{\mathrm{T}} \tilde{\boldsymbol{X}}\right)^{-1} \tilde{\boldsymbol{X}}^{\mathrm{T}} \boldsymbol{y}
$$

## Practical concerns

## Bottleneck of computing the LMS solution

$$
\boldsymbol{w}=\left(\tilde{\boldsymbol{X}}^{\mathrm{T}} \tilde{\boldsymbol{X}}\right)^{-1} \tilde{\boldsymbol{X}} \boldsymbol{y}
$$

Matrix multiply of $\tilde{\boldsymbol{X}}^{\mathrm{T}} \tilde{\boldsymbol{X}} \in \mathbb{R}^{(\mathrm{D}+1) \times(\mathrm{D}+1)}$ Inverting the matrix $\tilde{\boldsymbol{X}}^{\mathrm{T}} \tilde{\boldsymbol{X}}$

Scalable methods

- Batch gradient descent
- Stochastic gradient descent


## Stochastic gradient descent

Widrow-Hoff rule: update parameters using one example at a time

- Initialize $\tilde{\boldsymbol{w}}$ to $\tilde{\boldsymbol{w}}^{(0)}$ (anything reasonable is fine); set $t=0$; choose $\eta>0$
- Loop until convergence
(1) random choose a training a sample $\boldsymbol{x}_{t}$
(2) Compute its contribution to the gradient

$$
\boldsymbol{g}_{t}=\left(\tilde{\boldsymbol{x}}_{t}^{\mathrm{T}} \tilde{\boldsymbol{w}}^{(t)}-y_{t}\right) \tilde{\boldsymbol{x}}_{t}
$$

(3) Update the parameters

$$
\tilde{\boldsymbol{w}}^{(t+1)}=\tilde{\boldsymbol{w}}^{(t)}-\eta \boldsymbol{g}_{t}
$$

## Stochastic gradient descent

Widrow-Hoff rule: update parameters using one example at a time

- Initialize $\tilde{\boldsymbol{w}}$ to $\tilde{\boldsymbol{w}}^{(0)}$ (anything reasonable is fine); set $t=0$; choose $\eta>0$
- Loop until convergence
(1) random choose a training a sample $\boldsymbol{x}_{t}$
(2) Compute its contribution to the gradient

$$
\boldsymbol{g}_{t}=\left(\tilde{\boldsymbol{x}}_{t}^{\mathrm{T}} \tilde{\boldsymbol{w}}^{(t)}-y_{t}\right) \tilde{\boldsymbol{x}}_{t}
$$

(3) Update the parameters

$$
\tilde{\boldsymbol{w}}^{(t+1)}=\tilde{\boldsymbol{w}}^{(t)}-\eta \boldsymbol{g}_{t}
$$

(3) $t \leftarrow t+1$

## Stochastic gradient descent

Widrow-Hoff rule: update parameters using one example at a time

- Initialize $\tilde{\boldsymbol{w}}$ to $\tilde{\boldsymbol{w}}^{(0)}$ (anything reasonable is fine); set $t=0$; choose $\eta>0$
- Loop until convergence
(1) random choose a training a sample $\boldsymbol{x}_{t}$
(2) Compute its contribution to the gradient

$$
\boldsymbol{g}_{t}=\left(\tilde{\boldsymbol{x}}_{t}^{\mathrm{T}} \tilde{\boldsymbol{w}}^{(t)}-y_{t}\right) \tilde{\boldsymbol{x}}_{t}
$$

(3) Update the parameters

$$
\tilde{\boldsymbol{w}}^{(t+1)}=\tilde{\boldsymbol{w}}^{(t)}-\eta \boldsymbol{g}_{t}
$$

(3) $t \leftarrow t+1$

How does the complexity per iteration compare with gradient descent?

- $O(\mathrm{ND})$ for gradient descent versus $O(\mathrm{D})$ for SGD


## What if $\tilde{\boldsymbol{X}}^{\mathrm{T}} \tilde{\boldsymbol{X}}$ is not invertible

## Can you think of any reasons why that could happen?

Answer 1: $\mathrm{N}<\mathrm{D}$. Intuitively, not enough data to estimate all the parameters.

Answer 2: $\boldsymbol{X}$ columns are not linearly independent. Intuitively, there are two features that are perfectly correlated. In this case, solution is not unique.

## Ridge regression

For $\tilde{X}^{\mathrm{T}} \tilde{\boldsymbol{X}}$ that is not invertible

$$
\tilde{\boldsymbol{w}}=\left(\tilde{\boldsymbol{X}}^{\mathrm{T}} \tilde{\boldsymbol{X}}+\lambda \boldsymbol{I}\right)^{-1} \tilde{\boldsymbol{X}}^{\mathrm{T}} \boldsymbol{y}
$$

This is equivalent to adding an extra term to $R S S(\tilde{\boldsymbol{w}})$

$$
\overbrace{\frac{1}{2}\left\{\tilde{\boldsymbol{w}}^{\mathrm{T}} \tilde{\boldsymbol{X}}^{\mathrm{T}} \tilde{\boldsymbol{X}} \tilde{\boldsymbol{w}}-2\left(\tilde{\boldsymbol{X}}^{\mathrm{T}} \boldsymbol{y}\right)^{\mathrm{T}} \tilde{\boldsymbol{w}}\right\}}^{R S S(\tilde{\boldsymbol{w}})}+\underbrace{\frac{1}{2} \lambda\|\tilde{w}\|_{2}^{2}}_{\text {regularization }}
$$

What if data is not linearly separable or fits to a line Example of nonlinear classification


Example of nonlinear regression


## General nonlinear basis functions

## We can use a nonlinear mapping

$$
\boldsymbol{\phi}(\boldsymbol{x}): \boldsymbol{x} \in \mathbb{R}^{D} \rightarrow \boldsymbol{z} \in \mathbb{R}^{M}
$$

where $M$ is the dimensionality of the new feature/input $\boldsymbol{z}$ (or $\boldsymbol{\phi}(\boldsymbol{x})$ ). Note that $M$ could be either greater than $D$ or less than or the same.

With the new features, we can apply our learning techniques to minimize our errors on the transformed training data

- linear methods: prediction is based on $\boldsymbol{w}^{\mathrm{T}} \boldsymbol{\phi}(\boldsymbol{x})$
- other methods: nearest neighbors, decision trees, etc


## Regression with nonlinear basis

## Residual sum squares

$$
\sum_{n}\left[\boldsymbol{w}^{\mathrm{T}} \boldsymbol{\phi}\left(\boldsymbol{x}_{n}\right)-y_{n}\right]^{2}
$$

where $\boldsymbol{w} \in \mathbb{R}^{M}$, the same dimensionality as the transformed features $\phi(\boldsymbol{x})$.
The LMS solution can be formulated with the new design matrix

$$
\boldsymbol{\Phi}=\left(\begin{array}{c}
\boldsymbol{\phi}\left(\boldsymbol{x}_{1}\right)^{\mathrm{T}} \\
\boldsymbol{\phi}\left(\boldsymbol{x}_{2}\right)^{\mathrm{T}} \\
\vdots \\
\boldsymbol{\phi}\left(\boldsymbol{x}_{N}\right)^{\mathrm{T}}
\end{array}\right) \in \mathbb{R}^{N \times M}, \quad \boldsymbol{w}^{\mathrm{LMS}}=\left(\boldsymbol{\Phi}^{\mathrm{T}} \boldsymbol{\Phi}\right)^{-1} \boldsymbol{\Phi}^{\mathrm{T}} \boldsymbol{y}
$$

## Example with regression <br> Polynomial basis functions

$$
\phi(x)=\left[\begin{array}{c}
1 \\
x \\
x^{2} \\
\vdots \\
x^{M}
\end{array}\right] \Rightarrow f(x)=w_{0}+\sum_{m=1}^{M} w_{m} x^{m}
$$

## Example with regression

## Polynomial basis functions

$$
\boldsymbol{\phi}(x)=\left[\begin{array}{c}
1 \\
x \\
x^{2} \\
\vdots \\
x^{M}
\end{array}\right] \Rightarrow f(x)=w_{0}+\sum_{m=1}^{M} w_{m} x^{m}
$$

Fitting samples from a sine function: underrfitting as $f(x)$ is too simple



## Adding high-order terms

$M=3$


## Adding high-order terms

$$
M=3
$$


$\mathbf{M}=\mathbf{9}$ : overfitting


More complex features lead to better results on the training data, but potentially worse results on new data, e.g., test data!

## Overfitting

Parameters for higher-order polynomials are very large

|  | $M=0$ | $M=1$ | $M=3$ | $M=9$ |
| :--- | ---: | ---: | ---: | ---: |
| $w_{0}$ | 0.19 | 0.82 | 0.31 | 0.35 |
| $w_{1}$ |  | -1.27 | 7.99 | 232.37 |
| $w_{2}$ |  |  | -25.43 | -5321.83 |
| $w_{3}$ |  |  | 17.37 | 48568.31 |
| $w_{4}$ |  |  |  | -231639.30 |
| $w_{5}$ |  |  |  | 640042.26 |
| $w_{6}$ |  |  |  | -1061800.52 |
| $w_{7}$ |  |  |  | 1042400.18 |
| $w_{8}$ |  |  |  | -557682.99 |
| $w_{9}$ |  |  |  | 125201.43 |

## Overfitting can be quite disastrous

Fitting the housing price data with $M=3$


Note that the price would goes to zero (or negative) if you buy bigger ones! This is called poor generalization/overfitting.

How might we prevent overfitting?

## Outline

## (1) Administration

(2) Review of last lecture
(3) Basic ideas to overcome overfitting

- Use more training data
- Regularization methods


## 4 Bias/Variance Analysis

## Use more training data to prevent over fitting

The more, the merrier


## Use more training data to prevent over fitting

The more, the merrier




## Use more training data to prevent over fitting

The more, the merrier




What if we do not have a lot of data?

## Regularization methods

Intuition: For a linear model for regression

$$
\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}
$$

we can try to identify 'simpler' models. But what does it mean for a model to be simple?

## Regularization methods

Intuition: For a linear model for regression

$$
\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}
$$

we can try to identify 'simpler' models. But what does it mean for a model to be simple?

Assumption We can place a prior on our weights, assuming that $w_{d}$ is centered around zero.

With this reasoning, we will interpret $\boldsymbol{w}$ as a random variable and we will use the observed data $\mathcal{D}$ to update our prior belief on $\boldsymbol{w}$

## Review: Probabilistic interpretation for LMS

## Review: Probabilistic interpretation for LMS

- LMS model: $Y=\boldsymbol{w}^{\top} \boldsymbol{X}+\eta$
- $\eta \sim N\left(0, \sigma_{0}^{2}\right)$ is a Gaussian random variable
- Thus, $Y \sim N\left(\boldsymbol{w}^{\top} \boldsymbol{X}, \sigma_{0}^{2}\right)$


## Review: Probabilistic interpretation for LMS

- LMS model: $Y=\boldsymbol{w}^{\top} \boldsymbol{X}+\eta$
- $\eta \sim N\left(0, \sigma_{0}^{2}\right)$ is a Gaussian random variable
- Thus, $Y \sim N\left(\boldsymbol{w}^{\top} \boldsymbol{X}, \sigma_{0}^{2}\right)$
- We assume that $\boldsymbol{w}$ is fixed (Frequentist interpretation)


## Review: Probabilistic interpretation for LMS

- LMS model: $Y=\boldsymbol{w}^{\top} \boldsymbol{X}+\eta$
- $\eta \sim N\left(0, \sigma_{0}^{2}\right)$ is a Gaussian random variable
- Thus, $Y \sim N\left(\boldsymbol{w}^{\top} \boldsymbol{X}, \sigma_{0}^{2}\right)$
- We assume that $\boldsymbol{w}$ is fixed (Frequentist interpretation)
- We define $p\left(y \mid \boldsymbol{x}, \boldsymbol{w}, \sigma_{0}^{2}\right)$ as the sampling distribution given fixed values for the parameters $\boldsymbol{w}, \sigma_{0}^{2}$


## Review: Probabilistic interpretation for LMS

- LMS model: $Y=\boldsymbol{w}^{\top} \boldsymbol{X}+\eta$
- $\eta \sim N\left(0, \sigma_{0}^{2}\right)$ is a Gaussian random variable
- Thus, $Y \sim N\left(\boldsymbol{w}^{\top} \boldsymbol{X}, \sigma_{0}^{2}\right)$
- We assume that $\boldsymbol{w}$ is fixed (Frequentist interpretation)
- We define $p\left(y \mid \boldsymbol{x}, \boldsymbol{w}, \sigma_{0}^{2}\right)$ as the sampling distribution given fixed values for the parameters $\boldsymbol{w}, \sigma_{0}^{2}$
- The likelihood function maps parameters to probabilities

$$
L: \boldsymbol{w}, \sigma_{0}^{2} \mapsto p\left(\boldsymbol{y} \mid \mathcal{D}, \boldsymbol{w}, \sigma_{0}^{2}\right)=\prod_{n} p\left(y_{n} \mid \boldsymbol{x}_{n}, \boldsymbol{w}, \sigma_{0}^{2}\right)
$$

## Review: Probabilistic interpretation for LMS

- LMS model: $Y=\boldsymbol{w}^{\top} \boldsymbol{X}+\eta$
- $\eta \sim N\left(0, \sigma_{0}^{2}\right)$ is a Gaussian random variable
- Thus, $Y \sim N\left(\boldsymbol{w}^{\top} \boldsymbol{X}, \sigma_{0}^{2}\right)$
- We assume that $\boldsymbol{w}$ is fixed (Frequentist interpretation)
- We define $p\left(y \mid \boldsymbol{x}, \boldsymbol{w}, \sigma_{0}^{2}\right)$ as the sampling distribution given fixed values for the parameters $\boldsymbol{w}, \sigma_{0}^{2}$
- The likelihood function maps parameters to probabilities

$$
L: \boldsymbol{w}, \sigma_{0}^{2} \mapsto p\left(\boldsymbol{y} \mid \mathcal{D}, \boldsymbol{w}, \sigma_{0}^{2}\right)=\prod_{n} p\left(y_{n} \mid \boldsymbol{x}_{n}, \boldsymbol{w}, \sigma_{0}^{2}\right)
$$

- Maximizing likelihood with respect to $\boldsymbol{w}$ minimizes RSS and yields the LMS solution:

$$
\boldsymbol{w}^{\mathrm{LMS}}=\boldsymbol{w}^{\mathrm{ML}}=\arg \max _{\boldsymbol{w}} L\left(\boldsymbol{w}, \sigma_{0}^{2}\right)
$$

## Probabilistic interpretation of Ridge Regression

## Probabilistic interpretation of Ridge Regression

- Ridge Regression model: $Y=\boldsymbol{w}^{\top} \boldsymbol{X}+\eta$
- $Y \sim N\left(\boldsymbol{w}^{\top} \boldsymbol{X}, \sigma_{0}^{2}\right)$ is a Gaussian random variable (as before)
- $w_{d} \sim N\left(0, \sigma^{2}\right)$ are i.i.d. Gaussian random variables (unlike before)
- Note that all $w_{d}$ share the same variance $\sigma^{2}$


## Probabilistic interpretation of Ridge Regression

- Ridge Regression model: $Y=\boldsymbol{w}^{\top} \boldsymbol{X}+\eta$
- $Y \sim N\left(\boldsymbol{w}^{\top} \boldsymbol{X}, \sigma_{0}^{2}\right)$ is a Gaussian random variable (as before)
- $w_{d} \sim N\left(0, \sigma^{2}\right)$ are i.i.d. Gaussian random variables (unlike before)
- Note that all $w_{d}$ share the same variance $\sigma^{2}$
- $\boldsymbol{w}$ is a random variable with a prior distribution (Bayesian interpretation)


## Probabilistic interpretation of Ridge Regression

- Ridge Regression model: $Y=\boldsymbol{w}^{\top} \boldsymbol{X}+\eta$
- $Y \sim N\left(\boldsymbol{w}^{\top} \boldsymbol{X}, \sigma_{0}^{2}\right)$ is a Gaussian random variable (as before)
- $w_{d} \sim N\left(0, \sigma^{2}\right)$ are i.i.d. Gaussian random variables (unlike before)
- Note that all $w_{d}$ share the same variance $\sigma^{2}$
- $\boldsymbol{w}$ is a random variable with a prior distribution (Bayesian interpretation)
- To find $\boldsymbol{w}$ given data $\mathcal{D}$, we can compute posterior distribution of $\boldsymbol{w}$ :

$$
p(\boldsymbol{w} \mid \mathcal{D})=\frac{p(\mathcal{D} \mid \boldsymbol{w}) p(\boldsymbol{w})}{p(\mathcal{D})}
$$

## Probabilistic interpretation of Ridge Regression

- Ridge Regression model: $Y=\boldsymbol{w}^{\top} \boldsymbol{X}+\eta$
- $Y \sim N\left(\boldsymbol{w}^{\top} \boldsymbol{X}, \sigma_{0}^{2}\right)$ is a Gaussian random variable (as before)
- $w_{d} \sim N\left(0, \sigma^{2}\right)$ are i.i.d. Gaussian random variables (unlike before)
- Note that all $w_{d}$ share the same variance $\sigma^{2}$
- $\boldsymbol{w}$ is a random variable with a prior distribution (Bayesian interpretation)
- To find $\boldsymbol{w}$ given data $\mathcal{D}$, we can compute posterior distribution of $\boldsymbol{w}$ :

$$
p(\boldsymbol{w} \mid \mathcal{D})=\frac{p(\mathcal{D} \mid \boldsymbol{w}) p(\boldsymbol{w})}{p(\mathcal{D})}
$$

- Maximum a posterior (MAP) estimate:

$$
\boldsymbol{w}^{\mathrm{MAP}}=\arg \max _{\boldsymbol{w}} p(\boldsymbol{w} \mid \mathcal{D})=\arg \max _{\boldsymbol{w}} p(\mathcal{D}, \boldsymbol{w})
$$

## Probabilistic interpretation of Ridge Regression

- Ridge Regression model: $Y=\boldsymbol{w}^{\top} \boldsymbol{X}+\eta$
- $Y \sim N\left(\boldsymbol{w}^{\top} \boldsymbol{X}, \sigma_{0}^{2}\right)$ is a Gaussian random variable (as before)
- $w_{d} \sim N\left(0, \sigma^{2}\right)$ are i.i.d. Gaussian random variables (unlike before)
- Note that all $w_{d}$ share the same variance $\sigma^{2}$
- $\boldsymbol{w}$ is a random variable with a prior distribution (Bayesian interpretation)
- To find $\boldsymbol{w}$ given data $\mathcal{D}$, we can compute posterior distribution of $\boldsymbol{w}$ :

$$
p(\boldsymbol{w} \mid \mathcal{D})=\frac{p(\mathcal{D} \mid \boldsymbol{w}) p(\boldsymbol{w})}{p(\mathcal{D})}
$$

- Maximum a posterior (MAP) estimate:

$$
\boldsymbol{w}^{\mathrm{MAP}}=\arg \max _{\boldsymbol{w}} p(\boldsymbol{w} \mid \mathcal{D})=\arg \max _{\boldsymbol{w}} p(\mathcal{D}, \boldsymbol{w})
$$

- What's the relationship between MAP and MLE?


## Probabilistic interpretation of Ridge Regression

- Ridge Regression model: $Y=\boldsymbol{w}^{\top} \boldsymbol{X}+\eta$
- $Y \sim N\left(\boldsymbol{w}^{\top} \boldsymbol{X}, \sigma_{0}^{2}\right)$ is a Gaussian random variable (as before)
- $w_{d} \sim N\left(0, \sigma^{2}\right)$ are i.i.d. Gaussian random variables (unlike before)
- Note that all $w_{d}$ share the same variance $\sigma^{2}$
- $\boldsymbol{w}$ is a random variable with a prior distribution (Bayesian interpretation)
- To find $\boldsymbol{w}$ given data $\mathcal{D}$, we can compute posterior distribution of $\boldsymbol{w}$ :

$$
p(\boldsymbol{w} \mid \mathcal{D})=\frac{p(\mathcal{D} \mid \boldsymbol{w}) p(\boldsymbol{w})}{p(\mathcal{D})}
$$

- Maximum a posterior (MAP) estimate:

$$
\boldsymbol{w}^{\mathrm{MAP}}=\arg \max _{\boldsymbol{w}} p(\boldsymbol{w} \mid \mathcal{D})=\arg \max _{\boldsymbol{w}} p(\mathcal{D}, \boldsymbol{w})
$$

- What's the relationship between MAP and MLE?
- MAP reduces to MLE if we assume uniform prior for $p(\boldsymbol{w})$


## Estimating $\boldsymbol{w}$

- Let $X_{1}, \ldots, X_{N}$ be IID with $y \mid \boldsymbol{w}, \boldsymbol{x} \sim N\left(\boldsymbol{w}^{\top} \boldsymbol{x}, \sigma_{0}^{2}\right)$
- Let $w_{d}$ be IID with $w_{d} \sim N\left(0, \sigma^{2}\right)$

Joint likelihood of data and parameters (given $\sigma_{0}, \sigma$ )

$$
p(\mathcal{D}, \boldsymbol{w})=p(\mathcal{D} \mid \boldsymbol{w}) p(\boldsymbol{w})=\prod_{n} p\left(y_{n} \mid \boldsymbol{x}_{n}, \boldsymbol{w}\right) \prod_{d} p\left(w_{d}\right)
$$

## Estimating $\boldsymbol{w}$

- Let $X_{1}, \ldots, X_{N}$ be IID with $y \mid \boldsymbol{w}, \boldsymbol{x} \sim N\left(\boldsymbol{w}^{\top} \boldsymbol{x}, \sigma_{0}^{2}\right)$
- Let $w_{d}$ be IID with $w_{d} \sim N\left(0, \sigma^{2}\right)$

Joint likelihood of data and parameters (given $\sigma_{0}, \sigma$ )

$$
p(\mathcal{D}, \boldsymbol{w})=p(\mathcal{D} \mid \boldsymbol{w}) p(\boldsymbol{w})=\prod_{n} p\left(y_{n} \mid \boldsymbol{x}_{n}, \boldsymbol{w}\right) \prod_{d} p\left(w_{d}\right)
$$

Joint log likelihood Plugging in Gaussian PDF, we get:

$$
\begin{aligned}
\log p(\mathcal{D}, \boldsymbol{w}) & =\sum_{n} \log p\left(y_{n} \mid \boldsymbol{x}_{n}, \boldsymbol{w}\right)+\sum_{d} \log p\left(w_{d}\right) \\
& =-\frac{\sum_{n}\left(\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}_{n}-y_{n}\right)^{2}}{2 \sigma_{0}^{2}}-\sum_{d} \frac{1}{2 \sigma^{2}} w_{d}^{2}+\mathrm{const}
\end{aligned}
$$

## Estimating $\boldsymbol{w}$

- Let $X_{1}, \ldots, X_{N}$ be IID with $y \mid \boldsymbol{w}, \boldsymbol{x} \sim N\left(\boldsymbol{w}^{\top} \boldsymbol{x}, \sigma_{0}^{2}\right)$
- Let $w_{d}$ be IID with $w_{d} \sim N\left(0, \sigma^{2}\right)$

Joint likelihood of data and parameters (given $\sigma_{0}, \sigma$ )

$$
p(\mathcal{D}, \boldsymbol{w})=p(\mathcal{D} \mid \boldsymbol{w}) p(\boldsymbol{w})=\prod_{n} p\left(y_{n} \mid \boldsymbol{x}_{n}, \boldsymbol{w}\right) \prod_{d} p\left(w_{d}\right)
$$

Joint log likelihood Plugging in Gaussian PDF, we get:

$$
\begin{aligned}
\log p(\mathcal{D}, \boldsymbol{w}) & =\sum_{n} \log p\left(y_{n} \mid \boldsymbol{x}_{n}, \boldsymbol{w}\right)+\sum_{d} \log p\left(w_{d}\right) \\
& =-\frac{\sum_{n}\left(\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}_{n}-y_{n}\right)^{2}}{2 \sigma_{0}^{2}}-\sum_{d} \frac{1}{2 \sigma^{2}} w_{d}^{2}+\mathrm{const}
\end{aligned}
$$

MAP estimate: $\boldsymbol{w}^{\mathrm{MAP}}=\arg \max _{\boldsymbol{w}} \log p(\mathcal{D}, \boldsymbol{w})$

- As with LMS, set gradient equal to zero and solve (for $\boldsymbol{w}$ )


## Maximum a posterior (MAP) estimate

Regularized linear regression: a new error to minimize

$$
\mathcal{E}(\boldsymbol{w})=\sum_{n}\left(\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}_{n}-y_{n}\right)^{2}+\lambda\|\boldsymbol{w}\|_{2}^{2}
$$

where $\lambda>0$ is used to denote $\sigma_{0}^{2} / \sigma^{2}$. This extra term $\|\boldsymbol{w}\|_{2}^{2}$ is called regularization/regularizer and controls the model complexity.

## Maximum a posterior (MAP) estimate

Regularized linear regression: a new error to minimize

$$
\mathcal{E}(\boldsymbol{w})=\sum_{n}\left(\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}_{n}-y_{n}\right)^{2}+\lambda\|\boldsymbol{w}\|_{2}^{2}
$$

where $\lambda>0$ is used to denote $\sigma_{0}^{2} / \sigma^{2}$. This extra term $\|\boldsymbol{w}\|_{2}^{2}$ is called regularization/regularizer and controls the model complexity.

## Intuitions

## Maximum a posterior (MAP) estimate

Regularized linear regression: a new error to minimize

$$
\mathcal{E}(\boldsymbol{w})=\sum_{n}\left(\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}_{n}-y_{n}\right)^{2}+\lambda\|\boldsymbol{w}\|_{2}^{2}
$$

where $\lambda>0$ is used to denote $\sigma_{0}^{2} / \sigma^{2}$. This extra term $\|\boldsymbol{w}\|_{2}^{2}$ is called regularization/regularizer and controls the model complexity.

## Intuitions

- If $\lambda \rightarrow+\infty$, then $\sigma_{0}^{2} \gg \sigma^{2}$. That is, the variance of noise is far greater than what our prior model can allow for $\boldsymbol{w}$. In this case, our prior model on $\boldsymbol{w}$ would be more accurate than what data can tell us. Thus, we are getting a simple model. Numerically,

$$
\boldsymbol{w}^{\mathrm{MAP}} \rightarrow \mathbf{0}
$$

## Maximum a posterior (MAP) estimate

Regularized linear regression: a new error to minimize

$$
\mathcal{E}(\boldsymbol{w})=\sum_{n}\left(\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}_{n}-y_{n}\right)^{2}+\lambda\|\boldsymbol{w}\|_{2}^{2}
$$

where $\lambda>0$ is used to denote $\sigma_{0}^{2} / \sigma^{2}$. This extra term $\|\boldsymbol{w}\|_{2}^{2}$ is called regularization/regularizer and controls the model complexity.

## Intuitions

- If $\lambda \rightarrow+\infty$, then $\sigma_{0}^{2} \gg \sigma^{2}$. That is, the variance of noise is far greater than what our prior model can allow for $\boldsymbol{w}$. In this case, our prior model on $\boldsymbol{w}$ would be more accurate than what data can tell us. Thus, we are getting a simple model. Numerically,

$$
\boldsymbol{w}^{\mathrm{MAP}} \rightarrow \mathbf{0}
$$

- If $\lambda \rightarrow 0$, then we trust our data more. Numerically,

$$
\boldsymbol{w}^{\mathrm{MAP}} \rightarrow \boldsymbol{w}^{\mathrm{LMS}}=\arg \min \sum_{n}\left(\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}_{n}-y_{n}\right)^{2}
$$

## Closed-form solution

For regularized linear regression: the solution changes very little (in form) from the LMS solution

$$
\arg \min \sum_{n}\left(\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}_{n}-y_{n}\right)^{2}+\lambda\|\boldsymbol{w}\|_{2}^{2} \Rightarrow \boldsymbol{w}^{\mathrm{MAP}}=\left(\boldsymbol{X}^{\mathrm{T}} \boldsymbol{X}+\lambda \boldsymbol{I}\right)^{-1} \boldsymbol{X}^{\mathrm{T}} \boldsymbol{y}
$$

and reduces to the LMS solution when $\lambda=0$, as expected.

## Closed-form solution

For regularized linear regression: the solution changes very little (in form) from the LMS solution

$$
\arg \min \sum_{n}\left(\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}_{n}-y_{n}\right)^{2}+\lambda\|\boldsymbol{w}\|_{2}^{2} \Rightarrow \boldsymbol{w}^{\mathrm{MAP}}=\left(\boldsymbol{X}^{\mathrm{T}} \boldsymbol{X}+\lambda \boldsymbol{I}\right)^{-1} \boldsymbol{X}^{\mathrm{T}} \boldsymbol{y}
$$

and reduces to the LMS solution when $\lambda=0$, as expected.
If we have to use numerical procedure, the gradients and the Hessian matrix would change nominally too,

$$
\nabla \mathcal{E}(\boldsymbol{w})=2\left(\boldsymbol{X}^{\mathrm{T}} \boldsymbol{X} \boldsymbol{w}-\boldsymbol{X}^{\mathrm{T}} \boldsymbol{y}+\lambda \boldsymbol{w}\right), \quad \boldsymbol{H}=2\left(\boldsymbol{X}^{\mathrm{T}} \boldsymbol{X}+\lambda \boldsymbol{I}\right)
$$

As long as $\lambda \geq 0$, the optimization is convex.

## Example: fitting data with polynomials

## Our regression model

$$
y=\sum_{m=1}^{M} w_{m} x^{m}
$$

Regularization would discourage large parameter values as we saw with the LMS solution, thus potentially preventing overfitting.

|  | $M=0$ | $M=1$ | $M=3$ | $M=9$ |
| :--- | ---: | ---: | ---: | ---: |
| $w_{0}$ | 0.19 | 0.82 | 0.31 | 0.35 |
| $w_{1}$ |  | -1.27 | 7.99 | 232.37 |
| $w_{2}$ |  |  | -25.43 | -5321.83 |
| $w_{3}$ |  |  | 17.37 | 48568.31 |
| $w_{4}$ |  |  |  | -231639.30 |
| $w_{5}$ |  |  |  | 640042.26 |
| $w_{6}$ |  |  |  | -1061800.52 |
| $w_{7}$ |  |  |  | 1042400.18 |
| $w_{8}$ |  |  |  | -557682.99 |
| $w_{9}$ |  |  |  | 125201.43 |

## Overfitting in terms of $\lambda$

Overfitting is reduced from complex model to simpler one with the help of increasing regularizers




## Overfitting in terms of $\lambda$

Overfitting is reduced from complex model to simpler one with the help of increasing regularizers



$\lambda$ vs. residual error shows the difference of the model performance on training and testing dataset


## The effect of $\lambda$

Large $\lambda$ attenuates parameters towards 0

|  | $\ln \lambda=-\infty$ | $\ln \lambda=-18$ | $\ln \lambda=0$ |
| :--- | ---: | ---: | ---: |
| $w_{0}$ | 0.35 | 0.35 | 0.13 |
| $w_{1}$ | 232.37 | 4.74 | -0.05 |
| $w_{2}$ | -5321.83 | -0.77 | -0.06 |
| $w_{3}$ | 48568.31 | -31.97 | -0.06 |
| $w_{4}$ | -231639.30 | -3.89 | -0.03 |
| $w_{5}$ | 640042.26 | 55.28 | -0.02 |
| $w_{6}$ | -1061800.52 | 41.32 | -0.01 |
| $w_{7}$ | 1042400.18 | -45.95 | -0.00 |
| $w_{8}$ | -557682.99 | -91.53 | 0.00 |
| $w_{9}$ | 125201.43 | 72.68 | 0.01 |

## Regularized methods for classification

Adding regularizer to the cross-entropy functions used for binary and multinomial logistic regression

$$
\begin{aligned}
& \mathcal{E}(\boldsymbol{w})=-\sum_{n}\left\{y_{n} \log \sigma\left(\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}_{n}\right)+\left(1-y_{n}\right) \log \left[1-\sigma\left(\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}_{n}\right)\right]\right\}+\lambda\|\boldsymbol{w}\|_{2}^{2} \\
& \mathcal{E}\left(\boldsymbol{w}_{1}, \boldsymbol{w}_{2}, \ldots, \boldsymbol{w}_{K}\right)=-\sum_{n} \sum_{k} \log P\left(C_{k} \mid \boldsymbol{x}_{n}\right)+\lambda \sum_{k}\left\|\boldsymbol{w}_{k}\right\|_{2}^{2}
\end{aligned}
$$

## Regularized methods for classification

Adding regularizer to the cross-entropy functions used for binary and multinomial logistic regression

$$
\begin{aligned}
& \mathcal{E}(\boldsymbol{w})=-\sum_{n}\left\{y_{n} \log \sigma\left(\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}_{n}\right)+\left(1-y_{n}\right) \log \left[1-\sigma\left(\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}_{n}\right)\right]\right\}+\lambda\|\boldsymbol{w}\|_{2}^{2} \\
& \mathcal{E}\left(\boldsymbol{w}_{1}, \boldsymbol{w}_{2}, \ldots, \boldsymbol{w}_{K}\right)=-\sum_{n} \sum_{k} \log P\left(C_{k} \mid \boldsymbol{x}_{n}\right)+\lambda \sum_{k}\left\|\boldsymbol{w}_{k}\right\|_{2}^{2}
\end{aligned}
$$

## Numerical optimization

- Objective functions remain to be convex as long as $\lambda \geq 0$.
- Gradients and Hessians are changed marginally and can be easily derived.


## How to choose the right amount of regularization?

Can we tune $\lambda$ on the training dataset?

## How to choose the right amount of regularization?

Can we tune $\lambda$ on the training dataset?
No: as this will set $\lambda$ to zero, i.e., without regularization, defeating our intention to use it to control model complexity and to gain better generalization.
$\lambda$ is thus a hyperparmeter. To tune it,

- We can use a development/holdout dataset independent of training and testing dataset.
- We can do leave-one-out (LOO)

The procedure is similar to choose $K$ in the nearest neighbor classifiers.

## How to choose the right amount of regularization?

Can we tune $\lambda$ on the training dataset?
No: as this will set $\lambda$ to zero, i.e., without regularization, defeating our intention to use it to control model complexity and to gain better generalization.
$\lambda$ is thus a hyperparmeter. To tune it,

- We can use a development/holdout dataset independent of training and testing dataset.
- We can do leave-one-out (LOO)

The procedure is similar to choose $K$ in the nearest neighbor classifiers.
For different $\lambda$, we get $\boldsymbol{w}^{\mathrm{MAP}}$ and evaluate the model on the development/holdout dataset (or, the samples being left in LOO).

We then plot the curve $\lambda$ versus prediction error (accuracy, classification error) and find the place that the performance on the holdout/LOO is the best.

## Use cross-validation to choose $\lambda$

## Procedure

- Randomly partition training data into
$K$ disjoint parts
Normally, $K$ is chosen to be 10,5 , etc.
- For each possible value of $\lambda$
(1) Use one part as holdout; use other ( $K-1$ ) parts as training
(2) Evaluate the model on the holdout
(3) Do this $K$ times, and average the performance on the holdouts

- Choose the $\lambda$ with the best performance When $K=N$ (the number of training examples), this becomes LOO.


## Outline

(1) Administration
(2) Review of last lecture
(3) Basic ideas to overcome overfitting

4 Bias/Variance Analysis

## Basic and important machine learning concepts

## Supervised learning

We aim to build a function $h(\boldsymbol{x})$ to predict the true value $y$ associated with $\boldsymbol{x}$. If we make a mistake, we incur a loss

$$
\ell(h(\boldsymbol{x}), y)
$$

## Basic and important machine learning concepts

## Supervised learning

We aim to build a function $h(\boldsymbol{x})$ to predict the true value $y$ associated with $\boldsymbol{x}$. If we make a mistake, we incur a loss

$$
\ell(h(\boldsymbol{x}), y)
$$

Example: quadratic loss function for regression when $y$ is continuous

$$
\ell(h(\boldsymbol{x}), y)=[h(\boldsymbol{x})-y]^{2}
$$

Ex: when $y=0$


## Other types of loss functions

For classification: cross-entropy loss (also called logistic loss)
$\ell(h(\boldsymbol{x}), y)=-y \log h(\boldsymbol{x})-(1-y) \log [1-h(\boldsymbol{x})]$
Ex: when $y=1$


## Measure how good our predictor is

Risk: assume we know the true distribution of data $p(\boldsymbol{x}, y)$, the risk is

$$
R[h(\boldsymbol{x})]=\int_{\boldsymbol{x}, y} \ell(h(\boldsymbol{x}), y) p(\boldsymbol{x}, y) d \boldsymbol{x} d y
$$

## Measure how good our predictor is

Risk: assume we know the true distribution of data $p(\boldsymbol{x}, y)$, the risk is

$$
R[h(\boldsymbol{x})]=\int_{\boldsymbol{x}, y} \ell(h(\boldsymbol{x}), y) p(\boldsymbol{x}, y) d \boldsymbol{x} d y
$$

However, we cannot compute $R[h(\boldsymbol{x})]$, so we use empirical risk, given a training dataset $\mathcal{D}$

$$
R^{\mathrm{EMP}}[h(\boldsymbol{x})]=\frac{1}{N} \sum_{n} \ell\left(h\left(\boldsymbol{x}_{n}\right), y_{n}\right)
$$

## Measure how good our predictor is

Risk: assume we know the true distribution of data $p(\boldsymbol{x}, y)$, the risk is

$$
R[h(\boldsymbol{x})]=\int_{\boldsymbol{x}, y} \ell(h(\boldsymbol{x}), y) p(\boldsymbol{x}, y) d \boldsymbol{x} d y
$$

However, we cannot compute $R[h(\boldsymbol{x})]$, so we use empirical risk, given a training dataset $\mathcal{D}$

$$
R^{\mathrm{EMP}}[h(\boldsymbol{x})]=\frac{1}{N} \sum_{n} \ell\left(h\left(\boldsymbol{x}_{n}\right), y_{n}\right)
$$

Intuitively, as $N \rightarrow+\infty$,

$$
R^{\mathrm{EMP}}[h(\boldsymbol{x})] \rightarrow R[h(\boldsymbol{x})]
$$

## How this relates to what we have learned?

So far, we have been doing empirical risk minimization (ERM)

- For linear regression, $h(\boldsymbol{x})=\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}$, and we use squared loss
- For logistic regression, $h(\boldsymbol{x})=\sigma\left(\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}\right)$, and we use cross-entropy loss


## How this relates to what we have learned?

So far, we have been doing empirical risk minimization (ERM)

- For linear regression, $h(\boldsymbol{x})=\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}$, and we use squared loss
- For logistic regression, $h(\boldsymbol{x})=\sigma\left(\boldsymbol{w}^{\mathrm{T}} \boldsymbol{x}\right)$, and we use cross-entropy loss ERM might be problematic
- If $h(\boldsymbol{x})$ is complicated enough,

$$
R^{\mathrm{EMP}}[h(\boldsymbol{x})] \rightarrow 0
$$

- But then $h(\boldsymbol{x})$ is unlikely to do well in predicting things out of the training dataset $\mathcal{D}$
- This is called poor generalization or overfitting. We have just discussed approaches to address this issue.
- Let's try to understand why regularization might work from the context of the bias-variance tradeoff, focusing on regression / squared loss


## Bias/variance tradeoff (Looking ahead)

Error decomposes into 3 terms

$$
\mathbb{E}_{\mathcal{D}} R\left[h_{\mathcal{D}}(\boldsymbol{x})\right]=\text { VARIANCE }+\mathrm{BIAS}^{2}+\text { NOISE }
$$

We will prove this result, and interpret what it means...

