

CS145: INTRODUCTION TO DATA MINING

10: Vector Data: Density Estimation

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Methods Learnt: Last Lecture

	Vector Data	Set Data	Sequence Data	Text Data
Classification	Logistic Regression; Decision Tree; KNN SVM; NN			Naïve Bayes for Text
Clustering	K-means; hierarchical clustering; DBSCAN; Mixture Models			PLSA
Prediction	Linear Regression GLM*			
Frequent Pattern Mining		Apriori; FP growth	GSP; PrefixSpan	
Similarity Search			DTW	

Vector Data: Density Estimation

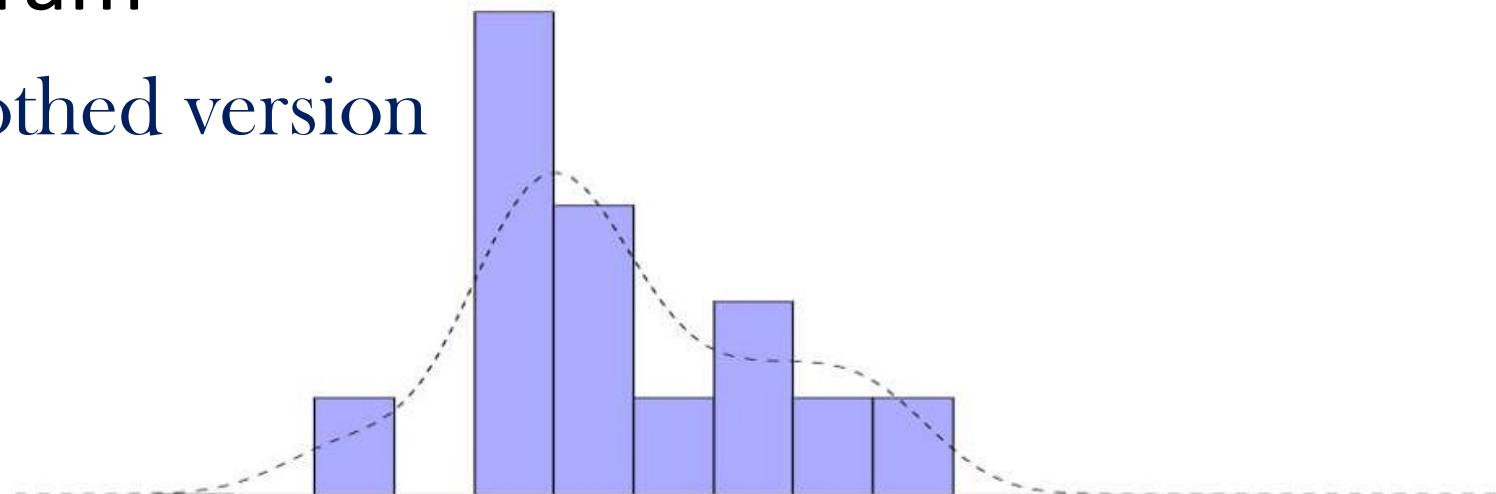
- Introduction 
- Nonparametric Density Estimation
- Parametric Density Estimation
- Summary

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Density Estimation from Data

- Goal
 - Estimate density function for a random variable from data
- Can be considered as an extension of histogram
 - Smoothed version



Recall

- Density-based clustering can be viewed as identifying connected dense areas of a distribution
- Critical for many other mining functions
 - Classification
 - Outlier detection

Nonparametric vs. parametric methods

- Nonparametric methods
 - No assumptions of the forms of the underlying densities
 - Can be used with arbitrary distributions
- Parametric methods
 - Have assumptions of the forms of the underlying densities
 - The densities are determined by fixed but unknown parameters

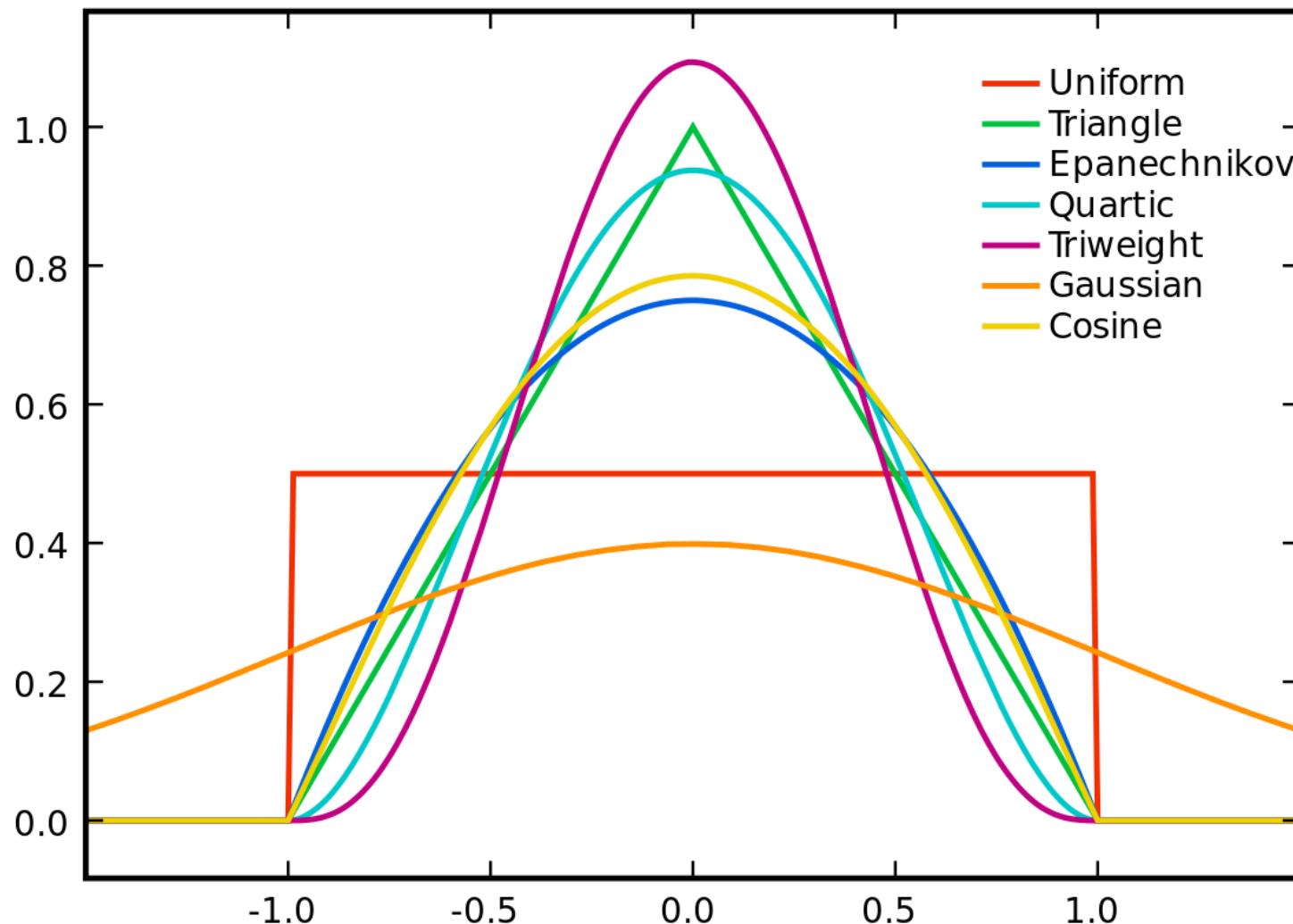
Vector Data: Density Estimation

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Kernel Density Estimation

- Given a dataset $D = (x_1, x_2, \dots, x_n)$, estimate its density function $f(x)$
- Kernel density estimator:
 - $\hat{f}_h(x) = \frac{1}{n} \sum_{i=1}^n K_h(x - x_i) = \frac{1}{nh} \sum_{i=1}^n K\left(\frac{x-x_i}{h}\right)$
 - h : bandwidth, controlling the smoothness of f
 - K : a non-negative real-valued integrable function, serving as weighting function
 - $\int_{-\infty}^{+\infty} K(u)du = 1$ (*normalization*)
 - $K(u) = K(-u)$ for all u (*symmetric*)

Examples of Kernels



Gaussian Kernel in 1-D case

- Example: Gaussian kernel

- $K(u) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}u^2}$

- Scaled kernel

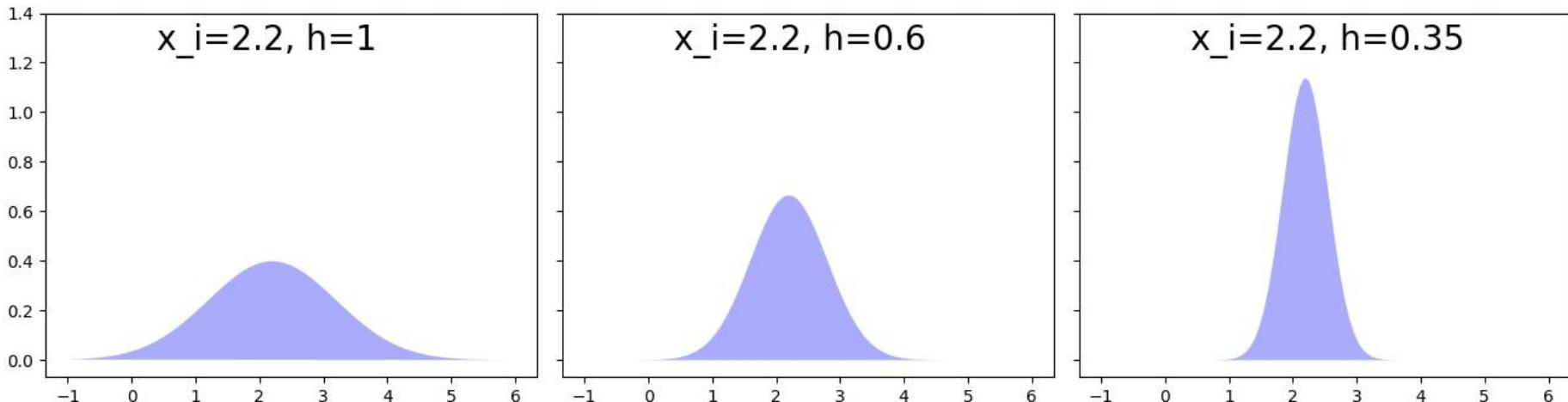
- $K_h(u) = \frac{1}{h} K\left(\frac{u}{h}\right)$

- In the Gaussian kernel case: $K_h(u) = \frac{1}{h\sqrt{2\pi}} e^{-\frac{u^2}{2h^2}}$

Influence from one data point

- The influence of x_i to x can be considered as a weighting function centered at x_i

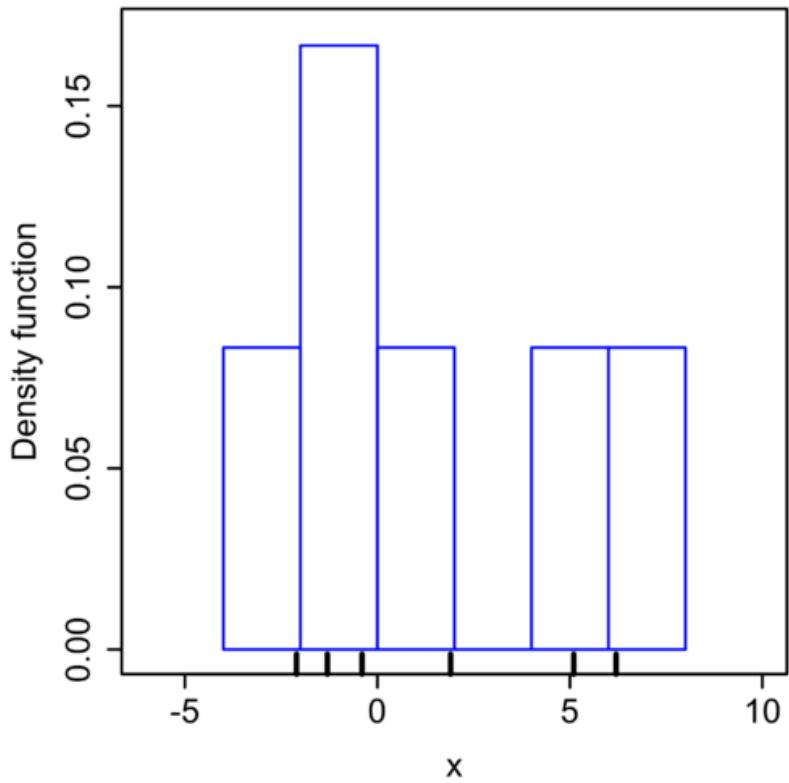
$$K_h(x - x_i) = \frac{1}{h\sqrt{2\pi}} e^{-\frac{(x-x_i)^2}{2h^2}}$$



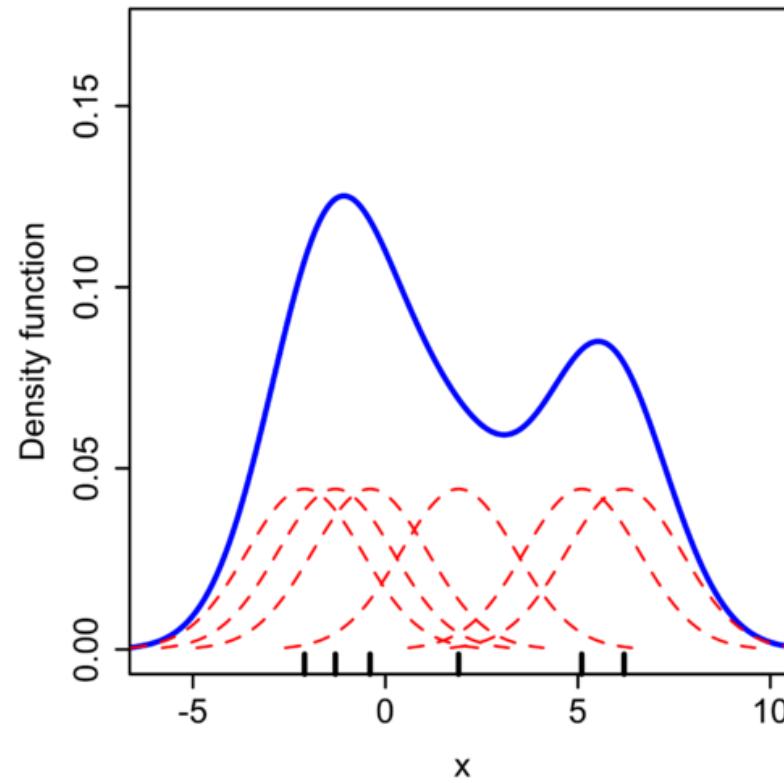
Recall: $\hat{f}_h(x) = \frac{1}{n} \sum_{i=1}^n K_h(x - x_i) = \frac{1}{nh} \sum_{i=1}^n K\left(\frac{x-x_i}{h}\right)$

Influence from multiple data points

- Aggregate influence from multiple data points to x



histogram



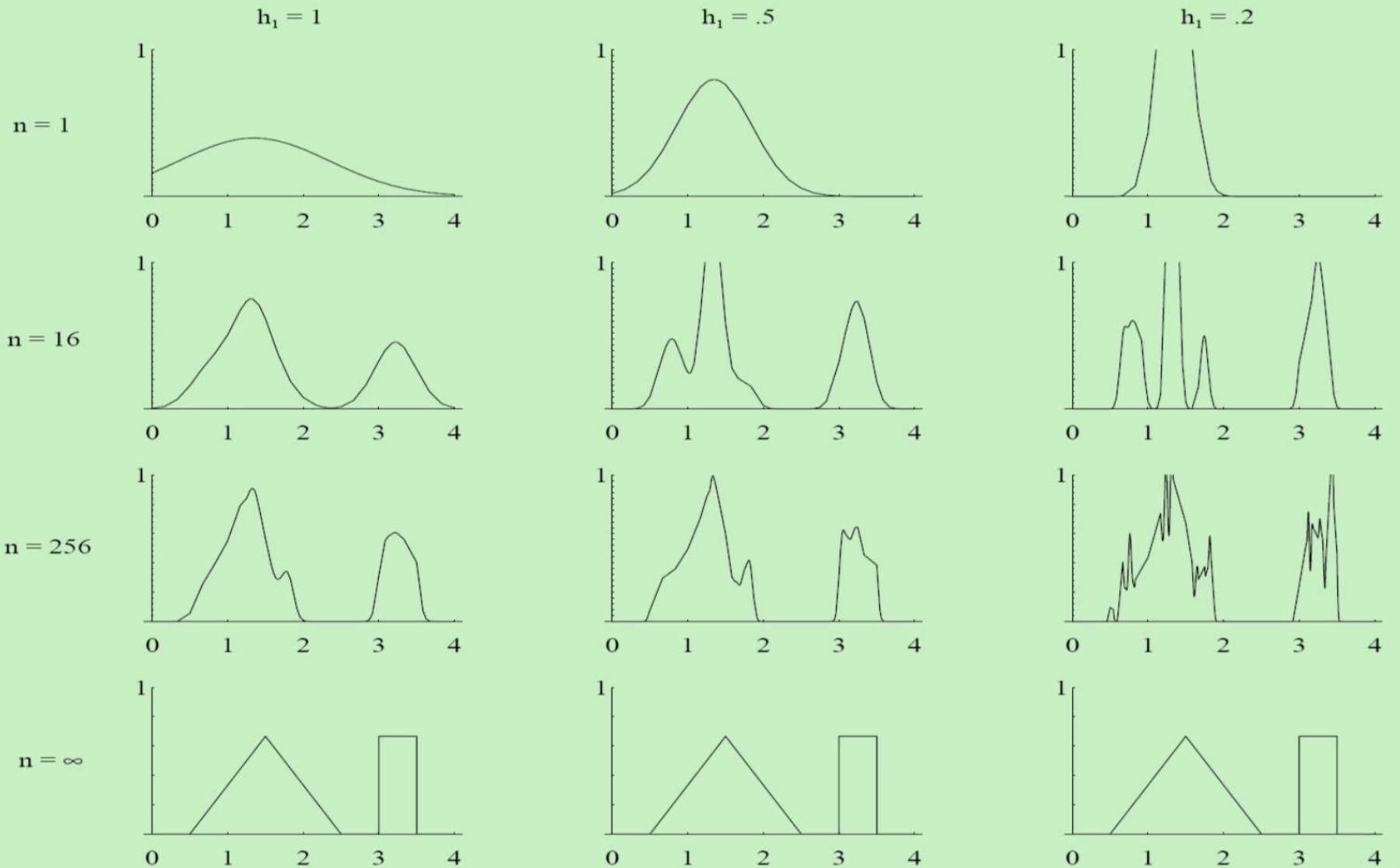
Density: Each red curve indicates $\frac{1}{n} K_h(x - x_i)$

Is it a density function?

- $\hat{f}_h(\mathbf{x}) = \frac{1}{n} \sum_{i=1}^n K_h(\mathbf{x} - \mathbf{x}_i)$
- A density function has to integrate to 1
- $K_h(\mathbf{x} - \mathbf{x}_i) = \frac{1}{h\sqrt{2\pi}} e^{-\frac{(x-x_i)^2}{2h^2}}$ integrates to 1
- Therefore, its average does so!

Impact of bandwidth

higher h , smoother density function



*DENCLUE: Using Statistical Density Functions for Clustering

- DENsity-based CLUstEring by Hinneburg & Keim (KDD'98)
- Using statistical density functions:

$$f_{Gaussian}(x, y) = e^{-\frac{d(x, y)^2}{2\sigma^2}}$$

influence of y on x

$$f_{Gaussian}^D(x) = \sum_{i=1}^N e^{-\frac{d(x, x_i)^2}{2\sigma^2}}$$

total influence
on x

$$\nabla f_{Gaussian}^D(x, x_i) = \sum_{i=1}^N (x_i - x) \cdot e^{-\frac{d(x, x_i)^2}{2\sigma^2}}$$

gradient of x in
the direction of x_i

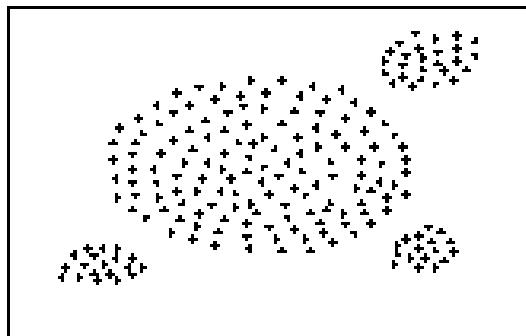
- Major features
 - Solid mathematical foundation
 - Good for data sets with large amounts of noise
 - Allows a compact mathematical description of arbitrarily shaped clusters in high-dimensional data sets
 - Significant faster than existing algorithm (e.g., DBSCAN)
 - But needs a large number of parameters

*Denclue: Technical Essence

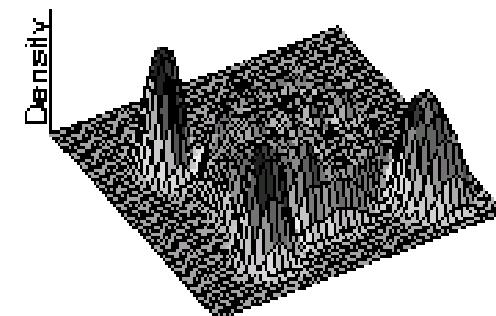
- Overall density of the data space can be calculated as the sum of the influence function of all data points
 - Influence function: describes the impact of a data point within its neighborhood
- Clusters can be determined mathematically by identifying density attractors
 - Density attractors are local maximal of the overall density function
 - Center defined clusters: assign to each density attractor the points density attracted to it
 - Arbitrary shaped cluster: merge density attractors that are connected through paths of high density ($>$ threshold)

*Density Attractor

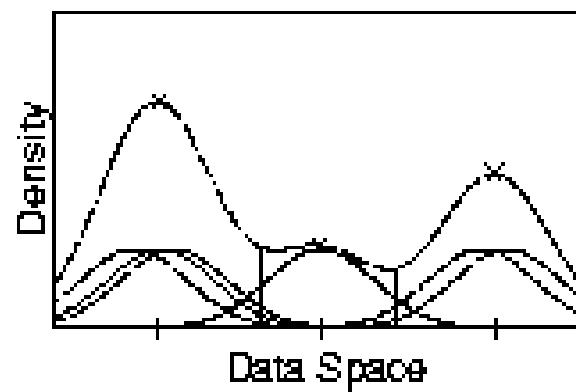
Can be detected by hill-climbing procedure of finding local maximums



(a) Data Set



(c) Gaussian



*Noise Threshold

- Noise Threshold ξ
 - Avoid trivial local maximum points
 - A point can be a density attractor only if
$$\hat{f}(x) \geq \xi$$

*Center-Defined and Arbitrary

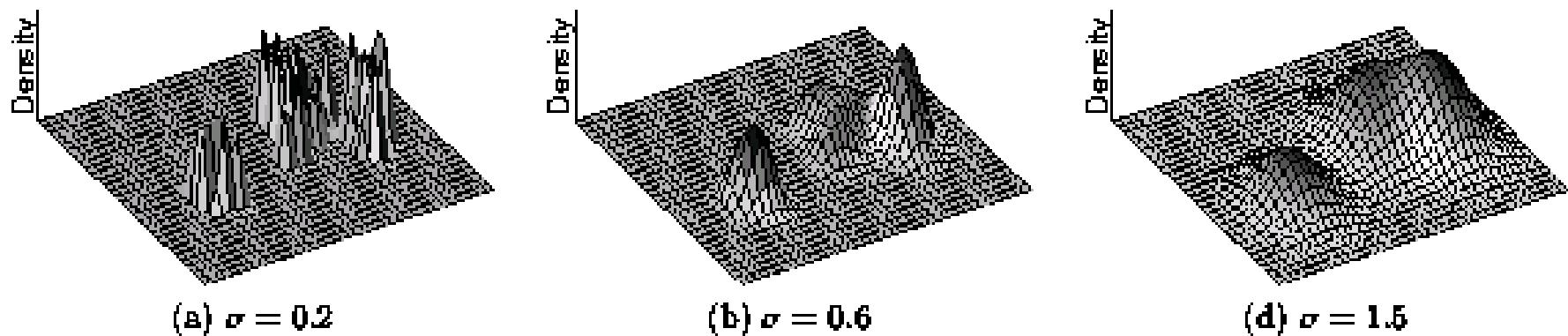


Figure 3: Example of Center-Defined Clusters for different σ

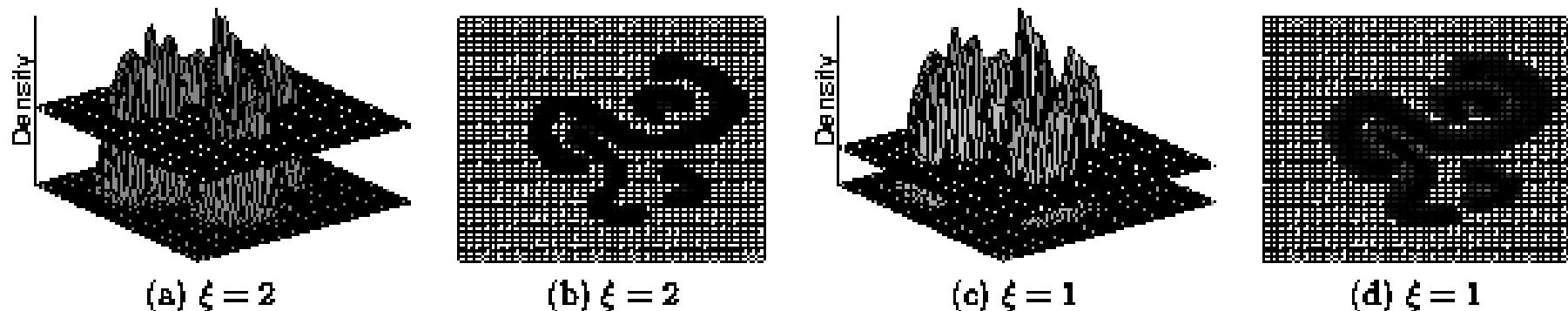


Figure 4: Example of Arbitrary-Shape Clusters for different ξ

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Maximum-Likelihood Estimation

- Data: $D = (x_1, x_2, \dots, x_n)$
- Parameters: θ
- Model: $p(x|\theta)$
- Likelihood of θ with respect to a set of data samples

$$L(\theta; D) = p(D|\theta) = \prod_{i=1}^n p(x_i|\theta)$$

- Maximum likelihood principle: find $\hat{\theta}$ that maximizes L
 - Agrees the most with the observation of current dataset

Log-likelihood function

- log-likelihood function

$$l(\boldsymbol{\theta}) \equiv \ln L(\boldsymbol{\theta}) = \ln p(D|\boldsymbol{\theta}) = \sum_i \ln p(x_i|\boldsymbol{\theta})$$

- Maximize likelihood function is equivalent to maximize log-likelihood function

$$\begin{aligned}\hat{\boldsymbol{\theta}} &= \arg \max_{\boldsymbol{\theta}} l(\boldsymbol{\theta}) \\ \Rightarrow \nabla_{\boldsymbol{\theta}} l(\boldsymbol{\theta}) &= 0\end{aligned}$$

$$\nabla_{\boldsymbol{\theta}} \equiv \begin{bmatrix} \frac{\partial}{\partial \theta_1} \\ \vdots \\ \frac{\partial}{\partial \theta_p} \end{bmatrix}$$

The Gaussian Case: Unknown Mean

- Consider 1-d Gaussian Distribution

$$x_i \sim N(\mu, \sigma^2)$$

where σ^2 is known, i.e., $\theta = \mu$

$$p(x_i | \mu) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

- The log-likelihood is then

$$l(\mu) = \sum_i \ln p(x_i | \mu) = \sum_i \left(-\frac{1}{2} \ln(2\pi\sigma^2) - \frac{(x_i - \mu)^2}{2\sigma^2} \right)$$

- The MLE estimator for μ is then

$$\nabla_{\mu} l(\mu) = 0 \Rightarrow \sum_i (x_i - \hat{\mu}) = 0 \Rightarrow \hat{\mu} = \frac{1}{n} \sum_i x_i$$

The Gaussian Case: Unknown Mean and Variance

- Consider 1-d Gaussian Distribution

$$x_i \sim N(\mu, \sigma^2)$$

where both μ and σ^2 are unknown, i.e., $\theta = (\mu, \sigma^2)$

$$p(x_i | \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x_i - \mu)^2}{2\sigma^2}}$$

- The log-likelihood is then

$$l(\mu, \sigma^2) = \sum_i \ln p(x_i | \mu, \sigma^2) = \sum_i \left(-\frac{1}{2} \ln(2\pi\sigma^2) - \frac{(x_i - \mu)^2}{2\sigma^2} \right)$$

- The MLE estimators for μ and σ^2 are then

$$\bullet \frac{\partial l(\mu, \sigma^2)}{\partial \mu} = 0 \Rightarrow \sum_i (x_i - \hat{\mu}) / \sigma^2 = 0 \Rightarrow \hat{\mu} = \frac{1}{n} \sum_i x_i$$

$$\bullet \frac{\partial l(\mu, \sigma^2)}{\partial \sigma^2} = 0 \Rightarrow \sum_i \left(-\frac{1}{2\widehat{\sigma}^2} + \frac{(x_i - \hat{\mu})^2}{2(\widehat{\sigma}^2)^2} \right) = 0 \Rightarrow \widehat{\sigma}^2 = \boxed{\frac{1}{n} \sum_i (x_i - \hat{\mu})^2}$$

Note it is biased

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Summary

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 - Maximum likelihood estimation